

Producers rational inattention and price stickiness: an inflated ordered probit approach

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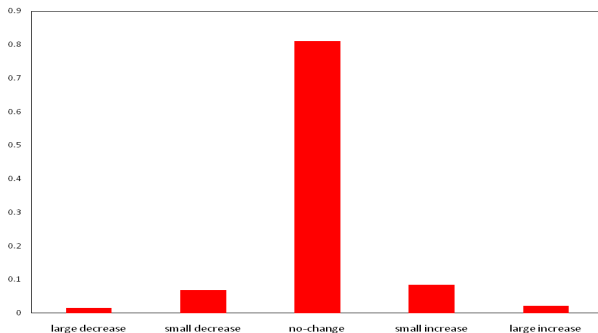
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Introduction and Background

- There are many papers (IPN and US) across many countries (US, Mexico, Hungary, *etc.*, *etc.*) looking at the presence of price inertia, or “stickiness” →
- ① Even with a changing economic environment, firms are reluctant to change prices → Banque de France Survey Responses



Introduction and Background

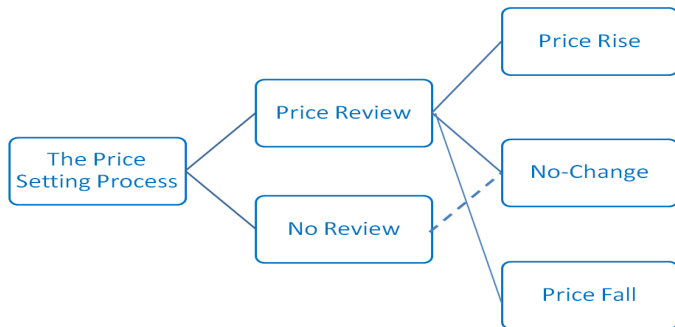
- The usual explanations for this inertia are menu costs and/or time-dependent behaviour
- 2. However, this does not account for the information we find in much survey data (see Fabiani *et al.*, 2006), which suggests that most (*manufacturing*) firms in the Euro area only review their prices very infrequently →
 - Fabiani *et al.*, (2006) - for the Euro area show that:
 - on average, only about 25% of firms review their prices at least monthly
 - this ranged from 10% (Italy, France, Austria, the Netherlands and Germany) to 30% (Belgium, Spain and Portugal)
 - however; nearly 60% of firms review their prices *at most* 3 times per year
 - ranged from 50% - 80%

Introduction and Background

- In the UK, *Hall et al.*, (1997): while only 30% of them declare a similar frequency for their price changes:
 - about 70% of UK firms review their prices at least quarterly
- 3. The majority of price reviews \rightarrow price changes
 - For the Euro area, only one price review in three \rightarrow price change:
 - two thirds of firms only change prices once a year at most
 - For the UK, only 30% of firms change their prices at least quarterly
 - So, in summary, we have:
 - infrequent price reviews, *plus*
 - a small conversion rate of these to price change, *equals*
 - sticky prices!

Introduction and Background

- The approach we suggest here, is to model price reviews and price changes simultaneously, and hence explain the high proportion of price inertia →



Price Reviews

- So, why do firms only periodically review their prices?
- Most obvious answer is cost!! (Sims, 1998, 2003)
- That is, there are significant costs in firms undertaking price reviews in order to determine their optimal prices →
- Zbaracki *et al.*, (2004) estimate these costs (net of consumer reactions and/or physical menu costs associated with the price change) for US manufacturing firm (with 8,000 products):
 - 30 man/months (\$250,000US at the time of the study!)
 - comprised of costs of gathering and processing the information needed for reviewing prices (around 11 man/months of labor: about \$100,000US);
 - and those of the decision making process itself (computation of new prices, simulation of alternative price strategies, and so on): 18 man/months: about \$150,000US
- These are big numbers, and represent a high proportion (1/4) of total costs incurred by a price change

Price Reviews

- So, likelihood of a price review will be a function of perceived costs relative to benefits (of the review, not necessarily a price change)
- However, these costs and expected benefits (of a price review) are unobserved
- Characteristics of the firm's environment should be good proxies here though:
 - a volatile environment \rightarrow \uparrow incentive for a review(s)
- However, even in a “stable” environment, there will also be factors that are likely to trigger a review
 - seasonality - many changes occur in January (\downarrow coordination problems due to the higher synchronization of price changes across firms and seasons: Konieczny and Rumler, 2005)
 - duration dependence - prices that do not change frequently appear to do so on a (broadly) yearly basis (may reflect the explicit/implicit contracts between firms and customers)

- We split the sources of potential volatility into the main three drivers of a firm's prices: production costs; product demand; and competitors' prices
 - “significant” movements in any of these are likely to trigger a review
- We also distinguish between two types of changes in the firm environment:
 - long run (“permanent shocks”) variations in price determinants; and
 - short-term (“transitory shocks”) variations

Price Changes

- Here, we base our approach on a standard state-dependent pricing model
- So, given that a price review has been undertaken, what are the likely factors that will trigger a subsequent change (or not)?
- Essentially, divergences of prevailing current price P_{it-1} from the optimal one P_{it}^* are likely to cause price changes (note, we have $i = 1, \dots, N$ firms observed over $t = 1, \dots, T_i$ periods)
- Assume monopolistic competition and a constant price elasticity of demand, given by a ($a < -1$), profit maximization leads to the usual equality:

- $$P_{it}^* = \frac{a}{1+a} MC_{it}, \text{ where } MC \text{ represents marginal cost}$$

Price Changes

- Assuming a simple static Cobb-Douglas cost function:



$$C_{it} = A_{ijt} Q_{it}^{\alpha} w_{it}^{\beta} \pi_{it}^{\gamma}$$

- Q_{it} = firm production level
 - w_{it} = represents the wage rate
 - π_{it} = the price of intermediate inputs
 - A_{ijt} = unobserved variables affecting costs, varying by sector j
- Then the 1st-order condition for output gives us an expression for MC which is substituted into the P_{it}^* equation giving:



$$P_{it}^* = \alpha \frac{a}{(a+1)} A_{ijt} Q_{it}^{\alpha-1} w_{it}^{\beta} \pi_{it}^{\gamma}, \text{ or in logs}$$

$$p_{it}^* = \ln \left[\alpha \frac{a}{(a+1)} \right] + \ln(A_{ijt}) + \dots + \gamma \ln(\pi_{it})$$

- We assume that A_{ijt} can be decomposed into three (multiplicative components):
 - a firm specific effect A_i ; a sector-specific effect, B_j ; and a third term representing a sectoral (common) time-varying component of prices C_{jt}
- Due to the relative dimensions of T and J , we proxy C_{jt} by sectoral production price indices at the NACE2 level
 $(PPI_{jt}; C_{jt} = PPI_{jt}^{\delta})$

Related Literature: Price Changes

- Thus we have proxies for p_{it}^* but the difference $(p_{it}^* - p_{it-1})$ is still unobserved! How to proceed...?
- For a *price spell* (starting at t_0), where price hasn't changed, we have
-

$$p_{it}^* - p_{it-1} = p_{it}^* - p_{it_0}$$

- Where the difference on the LHS $(p_{it}^* - p_{it-1})$ is exactly what we're interested in *i.e.*, the desired price change (Δp_{it}^d) !
- Assuming that (as usual in state-dependent pricing models) firms fully adjust to the optimal price level (when indeed, they change prices), then we have, at the start of the spell that prices were optimal:

$$p_{it_0} = p_{it_0}^*$$

- So, the desired price change (Δp_{it}^d) can be written as
-

$$\Delta p_{it}^d = p_{it}^* - p_{it-1} = p_{it}^* - p_{it_0}^*$$

Related Literature: Price Changes

- From above, we have expressions for p_{it}^* so all we have to do is difference the RHS of this equation, giving

$$\begin{aligned}\Delta p_{it}^d &= p_{it}^* - p_{it_0}^* \\ &= \Delta_s \ln(A_{ijt}) \\ &\quad + (\alpha - 1)\Delta_s \ln Q_{it} \\ &\quad + \beta \Delta_s \ln w_{it} \\ &\quad + \gamma \Delta_s \ln \pi_{it} + u_{it}\end{aligned}$$

- Where $\Delta_s x$ represents the variation of x over the course of the spell
- However, following Loupias and Sevestre (2007), we consider a more flexible form than simply the cumulative change in the x 's and let the effect of these vary over a given price spell

- Thus our final estimated equation is based on

$$\Delta p_{it}^d = \delta \Delta_s \ln(PPI_{jt}) + (\alpha - 1) \Delta_s \ln Q_{it} + \beta \Delta_s \ln w_{it} + \gamma \Delta_s \ln \pi_{it} + u_{it}$$

- Our “general” version of this replaces $\Delta_s x$ with the individual distributed lag function of Δx
- The desired price change, depends on:
 - 1 Firm specific variables: current and lagged changes in wages and prices of intermediate goods; current and lagged changes in the demand being addressed to the firm and
 - 2 Sector-specific and macro variables: e.g., variation in the sectoral inflation (common industry price shocks); note we also include here macro variables such as dummies for the VAT change (April 2000) and the Euro cash change-over (2002)

The Price Change Rule

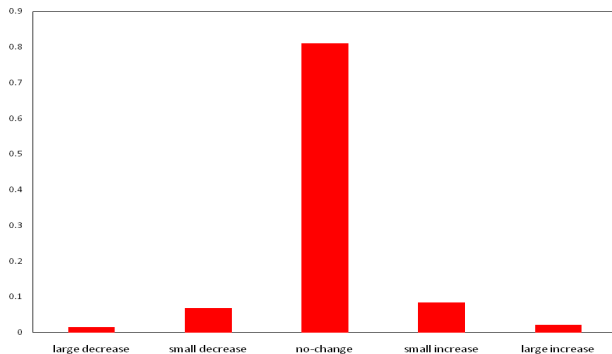
- The so-called (s, S) rule states when the foregone benefits of $P_{it} - P_{it}^*$ exceeds the costs, price is changed
- That is, when Δp_{it}^d is in excess of certain threshold values (μ) , observed prices will change
- Specifically, with $j = -2, \dots, 2$ outcomes observed in our data (“big” decreases to “big” increases: $J = 5$), we have

$$\Delta p_{it} = \begin{cases} -2 & \text{if } \Delta p_{it}^d \leq \mu_j, \\ j & \text{if } \mu_{j-1} < \Delta p_{it}^d \leq \mu_j, \quad j = -1, 0, 1 \\ 2 & \text{if } \mu_{J-1} \leq \Delta p_{it}^d \end{cases}$$

with μ_0 normalised to 0

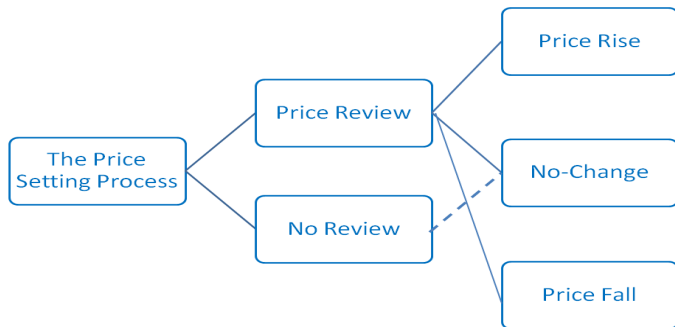
Econometric Model

- Thus our dependent variable, is an ordered discrete one → might suggest an ordered probit approach; but let's have a look at the raw data again...



Econometric Model

- *And*, we believe that the price-setting process can be decomposed into two sequential decisions →



- *But* we don't observe all price reviews

Econometric Model

- So, we want our model to explicitly allow for *jointly* the price review process and the price-rule/change process \rightarrow
- Let's start with a underlying latent variable, r_{it}^* , which is a firm's price review equation

$$r_{it}^* = \mathbf{x}'_{it}\boldsymbol{\beta} + u_{it}$$

- \mathbf{x} are our proxies for stability of the firm's environment *etc.*
- $r_{it}^* > 0 \rightarrow$ price review
- Under normality, the probability of this is, where $\Phi(\cdot)$ is the standard normal c.d.f.:

$$\Pr(r_{it} = 1 | \mathbf{x}_{it}) = \Pr(r_{it}^* > 0 | \mathbf{x}_{it}) = \Phi(\mathbf{x}'_{it}\boldsymbol{\beta})$$

- However, this needs to be combined with the price change process →
- Need to allow for price review firms to make a price decision, which may still be no change →
- Conditional on being in the price review regime, the price change process, Δp_{it}^d , kicks in...

$$\Delta p_{it}^d = \mathbf{z}'_{it} \boldsymbol{\gamma} + \varepsilon_{it}$$

- Where \mathbf{z} are the firms costs, demand etc., variables: denote this our *price change* equation

- *Conditional on $r_{it} = 1$* , probability of each “observed” Δp_{it} outcome (under normality) are

$$\Pr(\Delta p_{it}) = \begin{cases} \Pr(\Delta p_{it} = -1 | r_{it} = 1) = \Phi(\mu_1 - \mathbf{z}'_{it}\gamma) \\ \Pr(\Delta p_{it} = 0 | r_{it} = 1) = \Phi(\mu_{j-1} - \mathbf{z}'_{it}\gamma) - \\ \quad \Phi(\mu_j - \mathbf{z}'_{it}\gamma) \\ \Pr(\Delta p_{it} = 1 | r_{it} = 1) = 1 - \Phi(\mu_{j-2} - \mathbf{z}'_{it}\gamma) \end{cases}$$

- Under independence of ε and u the full probabilities for Δp_{it} , *unconditional on regime*, are (for $j = -1, 0, 1$)
- $\Pr(\Delta p_{it}) =$

$$\begin{cases} \Pr(\Delta p_{it} = -1) = \Phi(\mathbf{x}'_{it}\boldsymbol{\beta}) \Phi(\mu_0 - \mathbf{z}'_{it}\boldsymbol{\gamma}) \\ \Pr(\Delta p_{it} = 0) = [1 - \Phi(\mathbf{x}'_{it}\boldsymbol{\beta})] + \\ \quad \Phi(\mathbf{x}'_{it}\boldsymbol{\beta}) [\Phi(\mu_0 - \mathbf{z}'_{it}\boldsymbol{\gamma}) - \Phi(\mu_1 - \mathbf{z}'_{it}\boldsymbol{\gamma})] \\ \Pr(\Delta p_{it} = 1) = \Phi(\mathbf{x}'_{it}\boldsymbol{\beta}) [1 - \Phi(\mu_1 - \mathbf{z}'_{it}\boldsymbol{\gamma})] \end{cases}$$

- In this way, (along ZIP and ZIOP lines), the probability of a no-change outcome has been 'inflated'

Empirical Approach

- To observe a $\Delta p_{it} = 0$ (no-change) outcome we require either that:
 - $r_{it} = 0$ (the price review equation for no review dominates);
 - or jointly that $r_{it} = 1$ (review equation for review dominates) *and* that $0 < \Delta p_{it}^d \leq \mu$ (prevailing price not “far enough away” from optimal)
 - note observationally equivalent observations arise from two distinct sources!
- Can allow for a correlation between ε and u (equations relate to the same individual)
- Probabilities are now functions of the standardized bivariate normal c.d.f. with correlation coefficient $\rho_{\varepsilon u}$, $\Phi_2(a, b; \rho)$;
- $\Pr(\Delta p_{it}) =$

$$\begin{cases} \Pr(\Delta p_{it} = -1) = \Phi_2(\mathbf{x}'_{it}\boldsymbol{\beta}, -\mathbf{z}'_{it}\boldsymbol{\gamma}; -\rho_{\varepsilon u}) \\ \Pr(\Delta p_{it} = 0) = [1 - \Phi(\mathbf{x}'_{it}\boldsymbol{\beta})] + \begin{cases} \Phi_2(\mathbf{x}'_{it}\boldsymbol{\beta}, \mu - \mathbf{z}'_{it}\boldsymbol{\gamma}; -\rho_{\varepsilon u}) \\ -\Phi_2(\mathbf{x}'_{it}\boldsymbol{\beta}, -\mathbf{z}'_{it}\boldsymbol{\gamma}; -\rho_{\varepsilon u}) \end{cases} \\ \Pr(\Delta p_{it} = 1) = \Phi_2(\mathbf{x}'_{it}\boldsymbol{\beta}, \mathbf{z}'_{it}\boldsymbol{\gamma} - \mu; \rho_{\varepsilon u}) \end{cases}$$

Econometric Model

- We have panel data - can condition on unobserved firm heterogeneity in *both* equations \rightarrow

$$\begin{aligned}r_{it}^* &= \mathbf{x}'_{it}\boldsymbol{\beta} + \alpha_i + u_{it} \\ \Delta p_{it}^d &= \mathbf{z}'_{it}\boldsymbol{\gamma} + e_i + \varepsilon_{it}\end{aligned}$$

- Assume $\alpha_i \sim N(0, \sigma_\alpha^2)$ and $e_i \sim N(0, \sigma_e^2)$
- However, again, as these unobserved effects correspond to the same firm, correlations are likely \rightarrow

$$\begin{pmatrix} u_{it} \\ \varepsilon_{it} \end{pmatrix} \sim N \left[\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} \right]$$

- And

$$\begin{pmatrix} \alpha_i \\ e_i \end{pmatrix} \sim N \left[\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_\alpha^2 & \sigma_{\alpha e} \\ \sigma_{\alpha e} & \sigma_e^2 \end{pmatrix} \right]$$

Econometric Model

- Conditional on the individual effects, the (log-)likelihood, where $\theta = (\beta', \gamma', \mu, \rho, \sigma_\alpha^2, \sigma_e^2, \sigma_{\alpha e})'$
- $L(\theta | \alpha_i, e_i) =$

$$\sum_{i=1}^N \sum_{t=1}^{T_i} \sum_{j=0}^{J-1=2} d_{ijt} \ln [\Pr(\Delta p_{it} = j | \mathbf{x}_{it}, \mathbf{z}_{it})], \text{ where, } \Pr(\Delta p_{it})$$
$$= \begin{cases} \Phi_2(\mathbf{x}'_{it}\beta + \alpha_i, -\mathbf{z}'_{it}\gamma - e_i; -\rho_{\epsilon u}) \\ [1 - \Phi(\mathbf{x}'_{it}\beta + \alpha_i)] + \left\{ \begin{array}{l} \Phi_2(\mathbf{x}'_{it}\beta + \alpha_i, \mu - \mathbf{z}'_{it}\gamma - e_i; -\rho_{\epsilon u}) \\ -\Phi_2(\mathbf{x}'_{it}\beta + \alpha_i, -\mathbf{z}'_{it}\gamma - e_i; -\rho_{\epsilon u}) \end{array} \right\} \\ \Phi_2(\mathbf{x}'_{it}\beta + \alpha_i, \mathbf{z}'_{it}\gamma + e_i - \mu; \rho_{\epsilon u}), \end{cases}$$

- Thus estimation involves integration over both (α_i, e_i) bivariate normal integrals
- For estimation, need to remove the unobserved effects from these expressions \rightarrow

Empirical Approach

- Write the cholesky decomposition of Σ as

$$\text{chol}(\Sigma) = \text{chol} \begin{pmatrix} \sigma_{\alpha}^2 & \sigma_{\alpha e} \\ \sigma_{e\alpha} & \sigma_e^2 \end{pmatrix} = \begin{pmatrix} \delta_{11} & 0 \\ \delta_{12} & \delta_{22} \end{pmatrix}$$

- So that

$$\begin{pmatrix} \alpha_i \\ e_i \end{pmatrix} = \begin{pmatrix} \delta_{11} & 0 \\ \delta_{21} & \delta_{22} \end{pmatrix} \begin{pmatrix} v_{\alpha} \\ v_e \end{pmatrix}$$

- Where v_{α} and v_e are independent $N(0, 1)$ variables
- Now substitute α_i and e_i out using $\alpha_i = \delta_{11}v_{\alpha}$ and $e_i = \delta_{21}v_{\alpha} + \delta_{22}v_e$
- As v_{α} and v_e are just standard normal variates, they are most easily integrated out using simulation methods (*antithetic Halton draws*): simulated log-L is

$$L(\theta) = \sum_{i=1}^N \log \frac{1}{M} \sum_{m=1}^M \prod_{t=1}^{T_i} \sum_{j=0}^{J-1=2} d_{ijt} [\Pr(\Delta p_{it} = j | \mathbf{x}_{it}, \mathbf{z}_{it}, v_{\alpha}, v_e)]$$

Empirical Approach

- Marginal effects and probabilities can be evaluated by setting all random effects equal to their expected values (*i.e.*, zero) and similarly their covariance to zero
- Preferable approach: account for all random effects and correlations, and again use simulation methods
- Finally, we consider generalised Ordered Probit-type probabilities (Pudney and Shields, 2000) \rightarrow
- Allow boundary parameters μ , to be affected by firm characteristics, \mathbf{w}_{it} :

$$\mu_{ij} = \exp(\theta_j + \mathbf{w}'_{it}\boldsymbol{\phi})$$

- Collapses to the usual model if $\boldsymbol{\phi} = \mathbf{0}$

The dataset(s)

- Our database results from the merging of four different datasets:
 - ① The *Banque de France monthly business surveys* (1996-2005)
 - ② The *ACEMO survey* (French Ministry of Labour): includes information about *wages and employment* (quarterly, firm level, 1998-2005)
 - ③ *Monthly producer price indices computed by INSEE* at the 2-digit NACE level
 - ④ *Monthly industrial production indices computed by INSEE* (using the NES36 classification)
- Gives us *unbalanced panel of 42,954 observations with 5,019 firms/products*

The Dependent Variable

- The dependent variable is the answer to the question *By How Much Did Your Price Change Last Month?*
- Initially coded into 7 categories: large decrease up to large increase
- “Medium” and “large” responses, very sparse → re-coded to 5 outcomes

Price Review Equation, x

- Variables to account for the possible time-dependent timing of price reviews:
 - seasonal dummies ($month_j, j = 1$ to 11, except 8; ref. = December)
 - duration dummies ($dur1$ to $dur14$, ref. = duration of at least 15months)
- Variables accounting for the volatility (defined as inter-quartile ranges) of the environment:
 - firm specific volatility of costs (vi_wage, vi_iip)
 - volatility of firm production (vi_prod)
 - volatility of sector production (vi_ipi)
 - and volatility of competitors prices (vi_ipp)

3. Variables accounting for LR changes in the environment:

- 1 firm specific average growth rates of costs (ti_wage, ti_iip)
- 2 average growth rates of firm production (ti_prod)
- 3 average growth rates of sector production (ti_ipi)
- 4 and average growth rates of competitors prices (ti_ipp)
- 5 defined as 1 whenever a change in the variable is below the first quartile (strong negative shock) or above the third one (strong positive shock)

4. Variables accounting for SR changes in the environment:
 - ① dummies for transitory shocks in costs (si_wage , ti_iip)
 - ② dummies for transitory shocks in firm production (ti_prod)
 - ③ dummies for transitory shocks in sector production (ti_ipi)
 - ④ and dummies for transitory shocks in competitors prices (ti_ipp)
5. Industry and year dummies, plus dummies for specific events (euro-cash change-over and VAT rate change in April 2000)

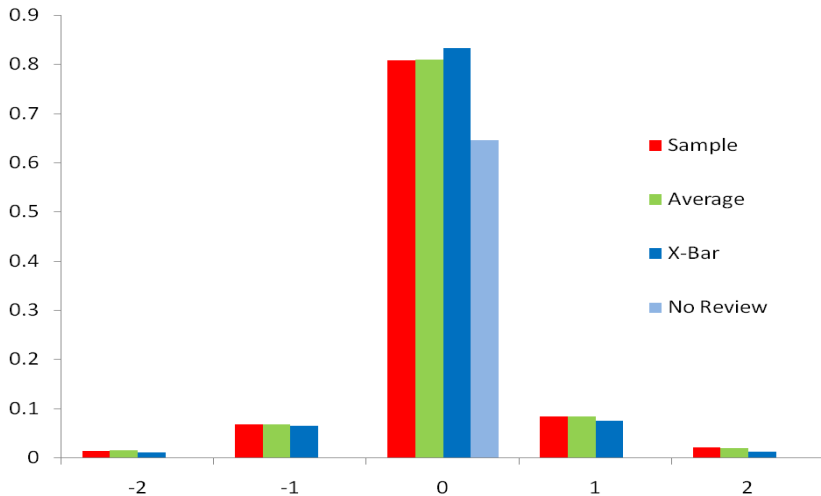
Price Change Equation, z

- 1 Variables accounting for observed variations in the firms environment:
 - 1 current and lagged changes in costs: wages ($wage$, $wage_l1$, etc.) and intermediate input prices (iip , iip_l1 , etc.)
 - 2 current and lagged changes in firm production ($prod$, $prod_l1$, etc.)
 - 3 and current and lagged changes in competitors prices (ipp , ipp_l1 , etc.)
- 2 Industry and year dummies, plus dummies for specific events (euro-cash change-over and VAT rate change in April 2000)

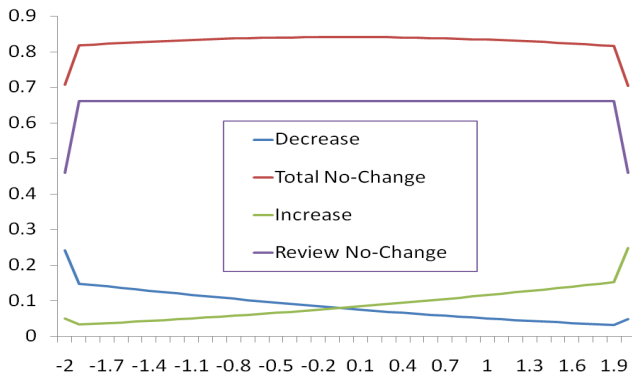
- There are several reasons to suspect some endogeneity in the regressors:
 - measurement errors (e.g., the timing of wage variations within quarters)
 - simultaneity of decisions regarding prices, production and wage changes
- So use IVs "à la Rivers-Vuong"
- Finally, to address the potential endogeneity associated with the implicit dynamic nature of the model, we also follow Wooldridge (2005) and include initial conditions in the model

- Still a work in progress!
- In general
 - evidence of random effects only in the review equation (differencing?)
 - small and negative correlation between the idiosyncratic errors
 - good significance levels across *both* equations

Overall Probabilities

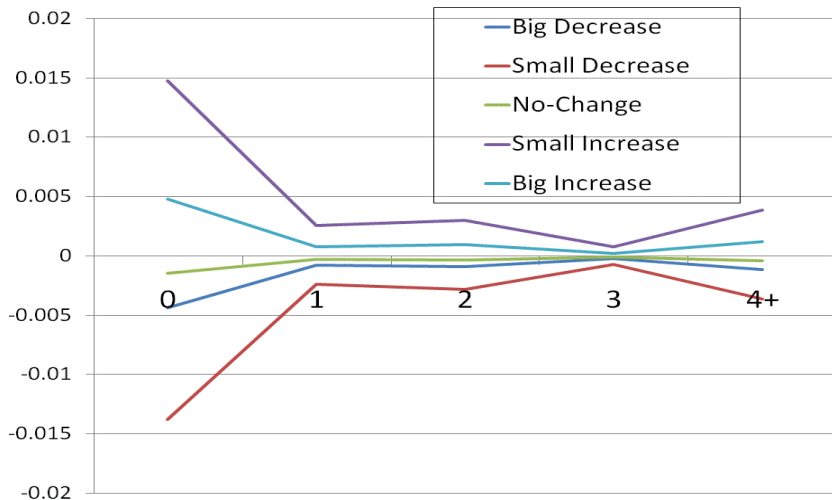


Raw Material Prices

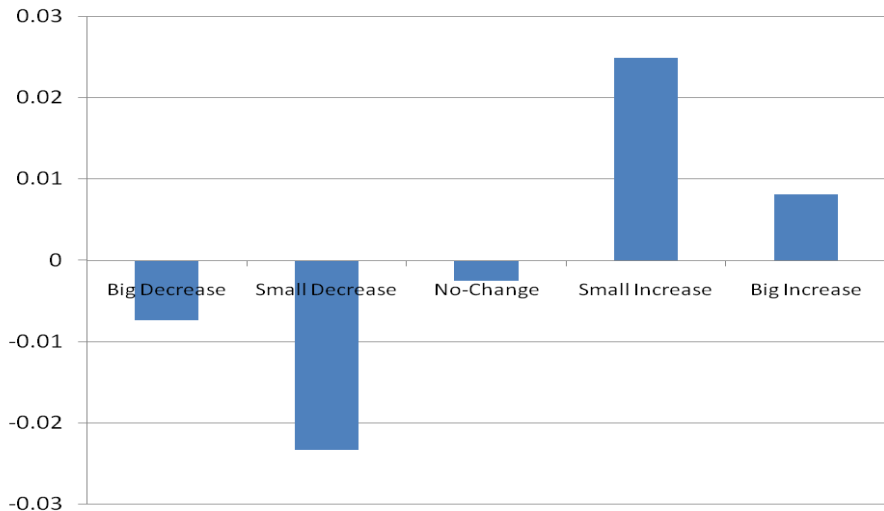


- No-change probabilities dominate
- Shock effects clear (but arbitrary)
- Increase (decrease) probabilities rise (fall) as prices raw material prices increase

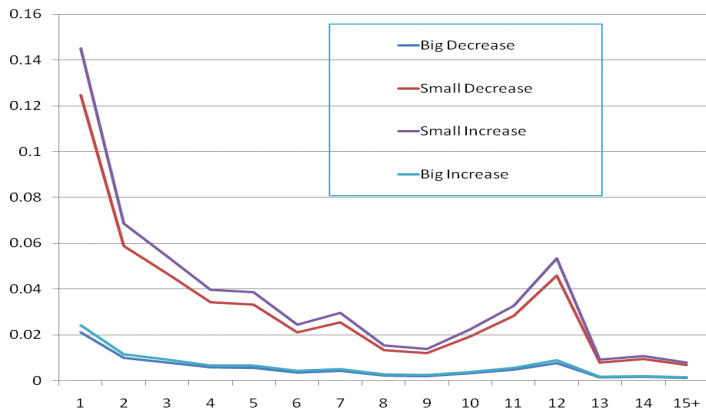
Marginal Effects: Wages Over Time



Total (LR) Wage Marginal Effects

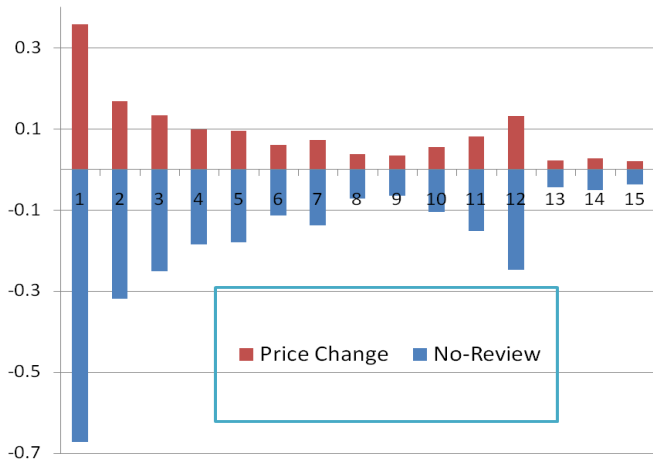


Duration Marginal Effects



- “Small” change probabilities uniformly dominate “large” ones
- The time-dependent peak at 12 months clearly evident
- Decreasing slopes, a result of heterogeneity across firms (Patrick and Hervé...)?

Duration Marginal Effects: No-Change Probabilities



- 1 *Price stickiness stems for a large part from the decision by firms not to review their prices on a continuous basis*
- 2 *Time-dependence is an important trigger of price reviews as well as shocks on intermediate input prices*
- 3 *Conditional on price reviews, prices react much more to changes in intermediate input costs than to changes in wages or demand (competitors' prices are also important)*
- 4 *However, the impact of wage changes on prices is far from negligible. It comes through the time-dependence component of the price-setting behavior*
 - 1 when we drop time-dependent variables in the price review equation → wages have a much more pronounced effect in both equations

5. *Policy implications?*

- 1 If you want to fully understand pricing behaviour, you need to explicitly take into account the price-review process and thus the inherent inertia in prices
- 2 Clear mix of time and state-behaviour in firm pricing behaviour
- 3 Increase number of price reviews → significant reduction in price stickiness - but how?! May be increase the number of shocks?!!

Merci Beaucoup, and Au Revoir!